

Markov Chains

Finite-State, Discrete-Time Markov Chains

Contents

- Concepts of Markov Process
- Discrete-Time Markov Chains
- Behavior of Markov Chains: Structural Properties

Concepts of Markov Process

- Markov Process
 - The future state depends only on the present state
- ?? Bernoulli process and Poisson process
- MP: Random Walk, Wiener Process, Sum Process, Random Telegraph
- Not MP: Moving Average Process
- Example: Customer arrival process!

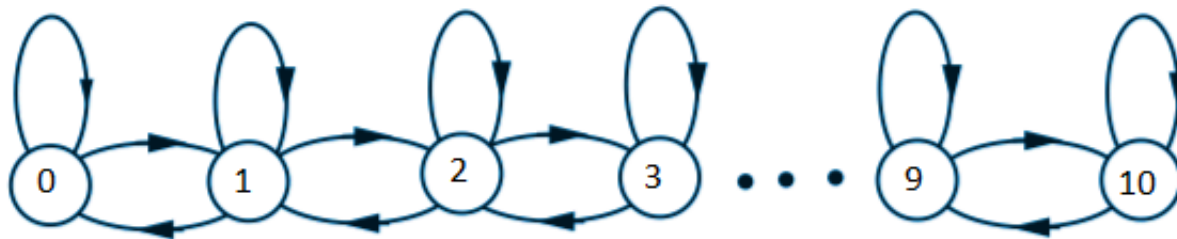
Concepts of Markov Process

- “State” X_n : Number of customers at time n .

Customer arrivals: Bernoulli(p)

geometric interarrival times

Customer service times: geometric(q)



Markov Process: Definition

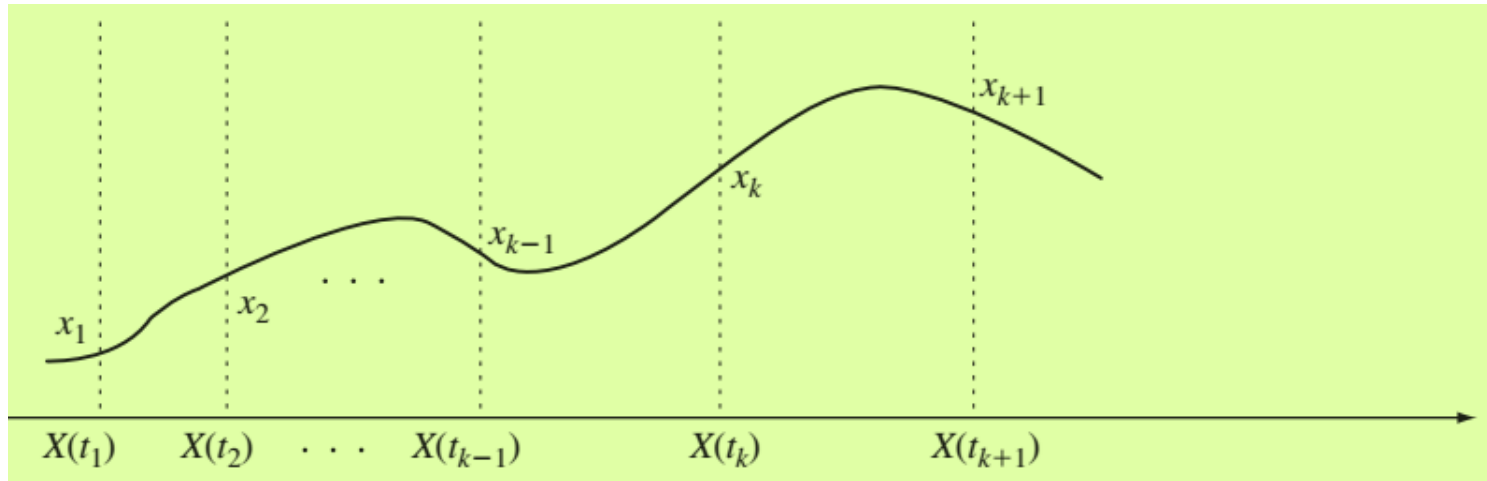
- If $X(t)$ is discrete-valued

$$\begin{aligned} P[X(t_{k+1}) = x_{k+1} \mid X(t_k) = x_k, \dots, X(t_1) = x_1] \\ = P[X(t_{k+1}) = x_{k+1} \mid X(t_k) = x_k] \end{aligned}$$

- If $X(t)$ is continuous-valued

$$\begin{aligned} P[a < X(t_{k+1}) \leq b \mid X(t_k) = x_k, \dots, X(t_1) = x_1] \\ = P[a < X(t_{k+1}) \leq b \mid X(t_k) = x_k] \end{aligned}$$

Markov Process: Definition



Markov property: Given $X(t_k)$, $X(t_{k+1})$ is independent of samples prior to t_k .

- If the samples of $X(t)$ are jointly continuous

$$f_{X(t_{k+1})}(x_{k+1} | X(t_k) = x_k, \dots, X(t_1) = x_1) = f_{X(t_{k+1})}(x_{k+1} | X(t_k) = x_k).$$

- State: pmf's and pdf's in a MP that are conditioned on several time instants always reduce to a pmf/pdf that is conditioned only on the most recent time instant.

Markov Chains

- An integer-valued Markov Process

If $X(t)$ is a Markov chain, then the joint pmf for three arbitrary time instants is

$$\begin{aligned} P[X(t_3) = x_3, X(t_2) = x_2, X(t_1) = x_1] \\ &= P[X(t_3) = x_3 | X(t_2) = x_2, X(t_1) = x_1] P[X(t_2) = x_2, X(t_1) = x_1] \\ &= P[X(t_3) = x_3 | X(t_2) = x_2] P[X(t_2) = x_2, X(t_1) = x_1] \\ &= P[X(t_3) = x_3 | X(t_2) = x_2] P[X(t_2) = x_2 | X(t_1) = x_1] P[X(t_1) = x_1], \end{aligned}$$

In General,

$$\begin{aligned} P[X(t_{k+1}) = x_{k+1}, X(t_k) = x_k, \dots, X(t_1) = x_1] \\ &= P[X(t_{k+1}) = x_{k+1} | X(t_k) = x_k] \end{aligned}$$

$$\begin{aligned} &P[X(t_k) = x_k | X(t_{k-1}) = x_{k-1}] \dots P[X(t_1) = x_1] \\ &= \left\{ \prod_{j=1}^k P[X(t_{j+1}) = x_{j+1} | X(t_j) = x_j] \right\} P[X(t_1) = x_1] \end{aligned}$$

Discrete-Time Markov Chains

Let X_n be a discrete-time integer-valued Markov chain that starts at $n = 0$ with pmf

$$p_j(0) \triangleq P[X_0 = j] \quad j = 0, 1, 2, \dots$$

- Assume
 - X_n is finite state
 - X_n takes on values on countable set of integers
- The joint pmf for the first $n+1$ values of the process

$$\begin{aligned} P[X_n = i_n, \dots, X_0 = i_0] \\ = P[X_n = i_n | X_{n-1} = i_{n-1}] \dots P[X_1 = i_1 | X_0 = i_0] P[X_0 = i_0]. \end{aligned}$$

Discrete-Time Markov Chains

- One-step transition probabilities

$$P[X_{n+1} = j | X_n = i] = p_{ij} \quad \text{for all } n.$$

Homogeneous

- So, the joint pmf for the first $n+1$ values of the process

$$P[X_n = i_n, \dots, X_0 = i_0] = p_{i_{n-1}, i_n} \dots p_{i_0, i_1} p_{i_0}(0).$$

Thus X_n is completely specified by the *initial pmf* $p_i(0)$ and the *matrix of one-step transition probabilities* P :

Transition probability matrix

$$P = \begin{bmatrix} p_{00} & p_{01} & p_{02} & \cdots \\ p_{10} & p_{11} & p_{12} & \cdots \\ \cdot & \cdot & \cdot & \cdot \\ p_{i0} & p_{i1} & \cdots & \cdot \\ \cdot & \cdot & \cdots & \cdot \end{bmatrix}.$$

$$1 = \sum_j P[X_{n+1} = j | X_n = i] = \sum_j p_{ij}.$$

Markov Chains: n-Step Transition Probabilities

$$p_{ij}(n) = P[X_{n+k} = j | X_k = i] \quad n \geq 0, i, j \geq 0.$$

- Two-step transition probabilities

$$\begin{aligned} P[X_2 = j, X_1 = k | X_0 = i] &= \frac{P[X_2 = j, X_1 = k, X_0 = i]}{P[X_0 = i]} \\ &= \frac{P[X_2 = j | X_1 = k]P[X_1 = k | X_0 = i]P[X_0 = i]}{P[X_0 = i]} \\ &= P[X_2 = j | X_1 = k]P[X_1 = k | X_0 = i] \\ &= p_{ik}(1)p_{kj}(1). \end{aligned}$$

Markov Chains: n-Step Transition Probabilities

$$p_{ij}(2) = \sum_k p_{ik}(1)p_{kj}(1) \quad \text{for all } i, j.$$

$$P(2) = P(1)P(1) = P^2.$$

Chapman–Kolmogorov equations:

$$p_{ij}(m + n) = \sum_k p_{ik}(m)p_{kj}(n) \quad \text{for all } n, m \geq 0 \text{ all } i, j.$$

$$P(n + m) = P(n)P(m).$$

$$P(n) = P^n.$$

Markov Chains: State Probabilities

- State probabilities at time n

$$\begin{aligned} p_j(n) &= \sum_i P[X_n = j | X_{n-1} = i] P[X_{n-1} = i] \\ &= \sum_i p_{ij} p_i(n-1). \end{aligned}$$

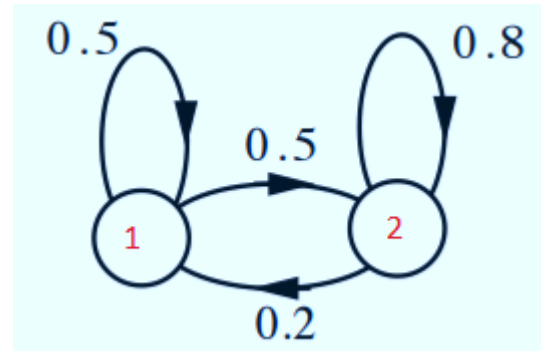
$$\mathbf{p}(n) = \mathbf{p}(n-1)P.$$

Similarly, $p_j(n)$ is related to $\mathbf{p}(0)$ by

$$\begin{aligned} p_j(n) &= \sum_i P[X_n = j | X_0 = i] P[X_0 = i] \\ &= \sum_i p_{ij}(n) p_i(0), \end{aligned}$$

$$\mathbf{p}(n) = \mathbf{p}(0)P^n = \mathbf{p}(0)P^n \quad n = 1, 2, \dots$$

Markov Chains: State Probabilities



	$n = 0$	$n = 1$	$n = 2$	$n = 100$	$n = 101$
$r_{11}(n)$	1	0.5	.35	$2/7$?
$r_{12}(n)$	0	0.5	.65	$5/7$?
$r_{21}(n)$	0	0.2			
$r_{22}(n)$	1	0.8			

Markov Chains: Steady State Probabilities

- As n approaches infinite $p_{ij}(n) \rightarrow \pi_j$ for all i .

$$P^n \rightarrow \mathbf{1}\pi$$

“Equilibrium” or “Steady state”

$$p_j(n) = \sum_i p_{ij}(n) p_i(0) \rightarrow \sum_i \pi_j p_i(0) = \pi_j.$$

$$\pi_j = \sum_i p_{ij} \pi_i,$$

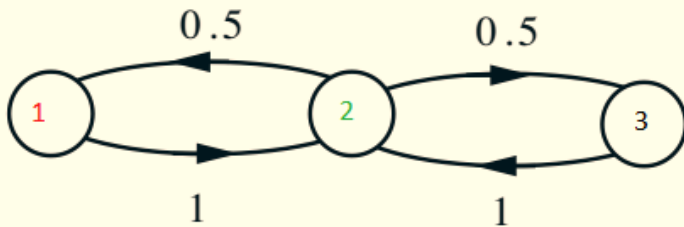
$$\pi = \pi P.$$

Underdetermined sys

$$\mathbf{p}(n) = \pi P^n = \pi \quad \text{for all } n.$$

Convergence of the Chain

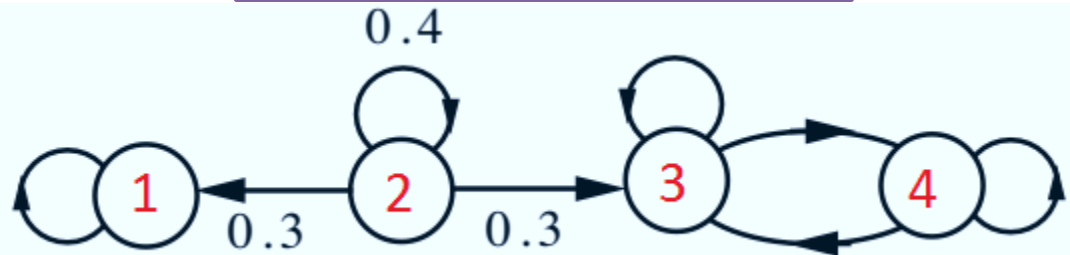
$P_{ij}(n)$ Converge?



n: even : $P_{22}=1$

n: odd: $P_{22}=0$

Does the Initial State matter?

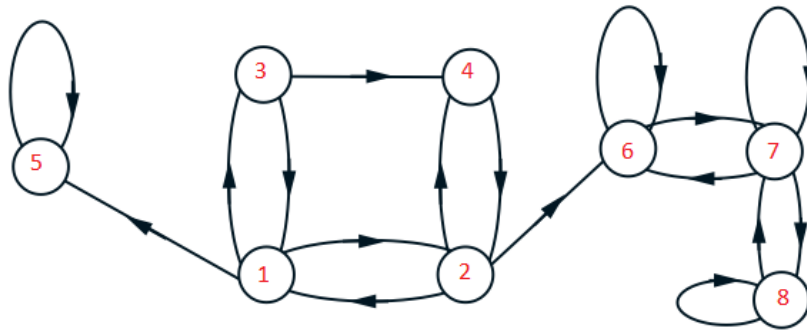


$P_{11}(n)=1$ for all n

$P_{31}(n)=0$ for all n

Classes of States

- The states of a Markov chain consist of one or more disjoint communication classes.
- A Markov chain that consists of a single class is said to be irreducible



State i is **recurrent** if:
starting from i ,
and from wherever you can go,
there is a way of returning to i

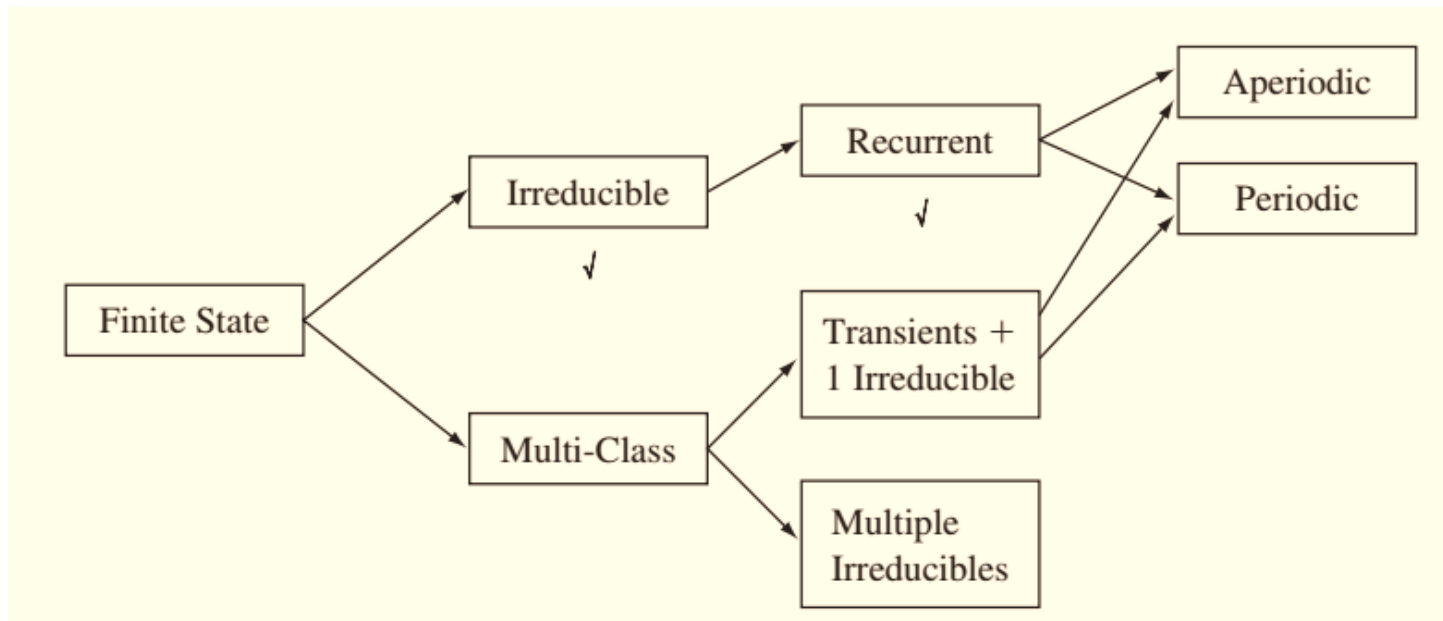
If not recurrent, called **transient**

$$f_i = P[\text{ever returning to state } i] = 1.$$

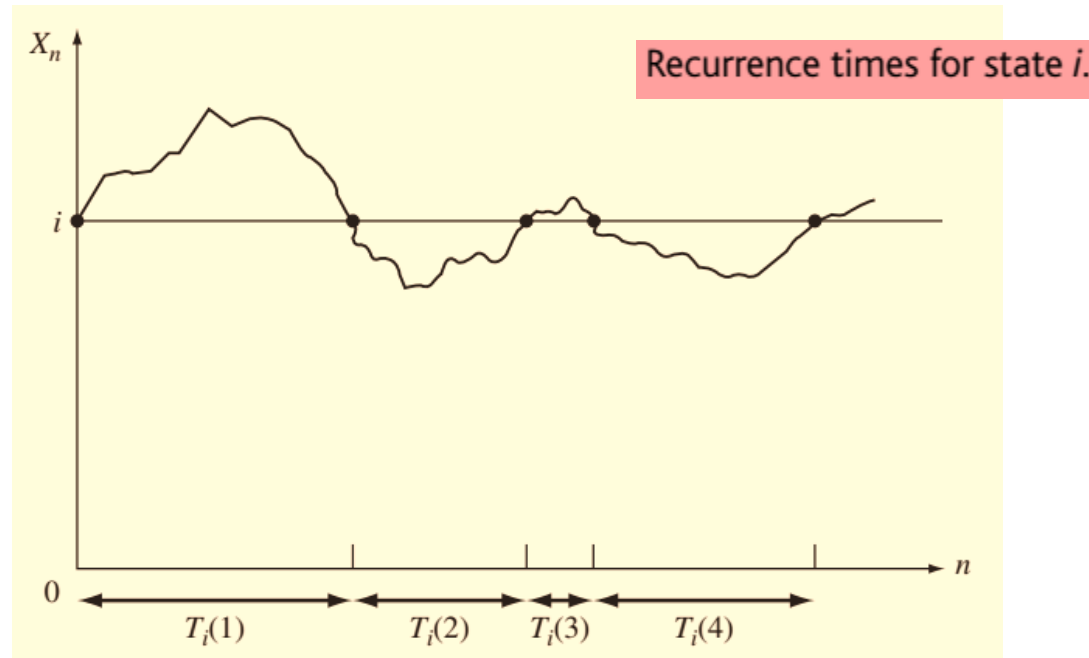
$$\sum_{n=1}^{\infty} p_{ii}(n) = \infty.$$

$$f_i < 1. \quad \sum_{n=1}^{\infty} p_{ii}(n) < \infty.$$

Finite State Markov Chain: structure



Limiting Probabilities



proportion of time in state $i = \frac{k}{T_i(1) + T_i(2) + \dots + T_i(k)}$.

proportion of time in state $i \rightarrow \frac{1}{E[T_i]} = \pi_i$,

If $E[T_i] < \infty$, then we say that state i is **positive recurrent**.

$\pi_i > 0$ if state i is positive recurrent.

If $E[T_i] = \infty$, then we say that state i is **null recurrent**.

$\pi_i = 0$ if state i is null recurrent.